

Computing expectations over p-boxes: two views (LP and RS) of the same problem

L. UTKIN¹ AND S. DESTERCKE²

¹State Forest Technical Academy, St. Petresburg, Russia ²IRSN (Institut de Radioprotection et de Sûreté Nucléaire), Cadarache, France



LIMSI

Problem statement

- information on r.v. X modeled by a p-box $[\underline{F}, \overline{F}]$
- lower $(\underline{\mathbb{E}})$ and upper $(\overline{\mathbb{E}})$ exp. on continuous function h(X):

$$\underline{\mathbb{E}}(h) = \inf_{\underline{F} \le F \le \overline{F}} \int_{\mathbb{R}}^{h(x)} \mathrm{d}F, \quad \overline{\mathbb{E}}(h) = \sup_{\underline{F} \le F \le \overline{F}} \int_{\mathbb{R}}^{h(x)} \mathrm{d}F$$
 (1)

Find optimal distribution F ($\underline{F}(x) \leq F(x) \leq \overline{F}(x)$) for which $\underline{\mathbb{E}}, \overline{\mathbb{E}}$ are reached.

Linear programming (LP) general view

Approximate F by N points $F(x_i)$, i = 1, ..., N and solve

$$\underline{\mathbb{E}}^*(h) = \inf \sum_{k=1}^N h(x_k) z_k$$
 or $\overline{\mathbb{E}}^*(h) = \sup \sum_{k=1}^N h(x_k) z_k$

subject to

$$z_i \ge 0, \quad i = 1, ..., N, \sum_{k=1}^{N} z_k = 1,$$

$$\sum_{k=1}^{i} z_k \le \overline{F}(x_i), \sum_{k=1}^{i} z_k \ge \underline{F}(x_i), \quad i = 1, ..., N.$$

- ightharpoonup $\underline{\mathbb{E}}^*, \overline{\mathbb{E}}^*$ are approximations of $\underline{\mathbb{E}}, \overline{\mathbb{E}}$
- ightharpoonup If N high, computation costs increase, and if N low, approximations can be bad ones

Random set (RS) general view

Mapping Γ from prob. space to power set $\wp(X)$ of a space X Here, mapping from [0,1] with Lebesgue measure to measurable subsets of \mathbb{R} .

Given continuous p-box $[\underline{F}, \overline{F}]$ $A_{\gamma} = [a_{*\gamma}, a_{\gamma}^*]$ is the set s.t.

$$a_{*\gamma} := \sup\{\overline{F}(x) < \gamma\} = \overline{F}^{-1}(\gamma),$$

 $a_{\gamma}^* := \inf\{\underline{F}(x) > \gamma\} = \underline{F}^{-1}(\gamma),$

P-box $[\underline{F}, \overline{F}]$ equivalent continuous random set with unif. density on [0,1] and

$$\Gamma(\gamma) = A_{\gamma} = [a_{*\gamma}, a_{\gamma}^*] \ \gamma \in [0, 1].$$

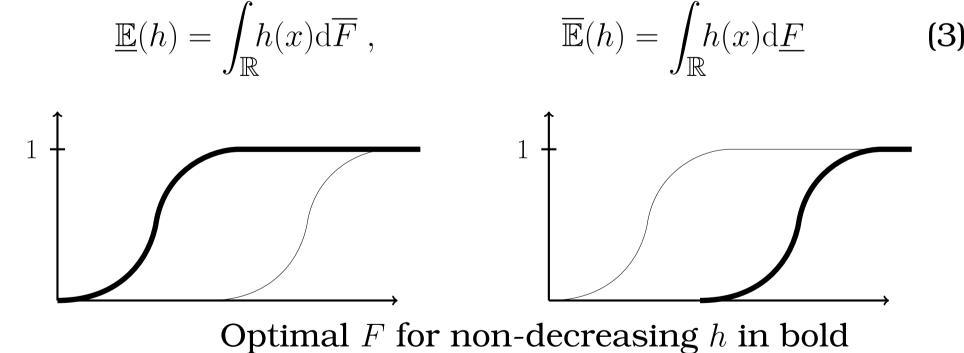
ightharpoonup Computing $\underline{\mathbb{E}},\overline{\mathbb{E}}$ of h can be reformulated

$$\underline{\mathbb{E}}(h) = \int_0^1 \inf_{x \in A_\gamma} h(x) \ d\gamma, \qquad \overline{\mathbb{E}}(h) = \int_0^1 \sup_{x \in A_\gamma} h(x) \ d\gamma. \qquad \textbf{(2)}$$

- Solution easily approximated by discretizing p-box on finite number of levels γ_i . Finding $\inf(\sup)$ on many levels can be difficult, and choosing too few γ_i or poor heuristics can again lead to bad approximations.
- For both approaches, need to find efficient AND reliable algorithms to compute $\underline{\mathbb{E}}, \overline{\mathbb{E}}$.
- \blacktriangleright Here, we interest ourselves to the case where h behavior is partially known

The easy case of monotonic functions

ightharpoonup If h is non-decreasing in \mathbb{R} , then we have :



One dimension, One maximum

▶ h has one maximum at point a and is increasing (decreasing) in $(-\infty, a]$ ($[a, \infty)$).

Unconditional expectations

▶ upper and lower expectations of h(X) on $[\underline{F}, \overline{F}]$ are

$$\overline{\mathbb{E}}(h) = \int_{-\infty}^{a} h(x) d\underline{F} + h(a) \left[\overline{F}(a) - \underline{F}(a) \right] + \int_{a}^{\infty} h(x) d\overline{F}$$

$$\underline{\mathbb{E}}(h) = \int_{-\infty}^{\overline{F}^{-1}(\alpha)} h(x) d\overline{F} + \int_{\underline{F}^{-1}(\alpha)}^{\infty} h(x) d\underline{F}$$
(5)

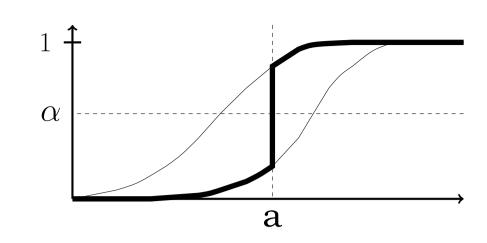
or, equivalently

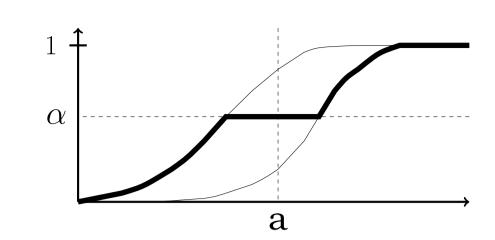
$$\overline{\mathbb{E}}(h) = \int_0^{\underline{F}(a)} h(a_{\gamma}^*) d\gamma + [\overline{F}(a) - \underline{F}(a)]h(a) + \int_{\overline{F}(a)}^1 h(a_{*\gamma}) d\gamma \qquad \textbf{(6)}$$

$$\underline{\mathbb{E}}(h) = \int_0^\alpha h(a_{*\gamma})d\gamma + \int_\alpha^1 h(a_\gamma^*)d\gamma, \qquad (7)$$

where α is one of the solution of the equation

$$h\left(\overline{F}^{-1}(\alpha)\right) = h\left(\underline{F}^{-1}(\alpha)\right). \tag{8}$$





Optimal F for $\overline{\mathbb{E}}(h)$ (vert. jump) Optimal F for $\underline{\mathbb{E}}(h)$ (hor. jump)

LP approach suggest to analytically find the level α , or to approximate solution by scanning different values of α .

Following formula derived with the RS approach

$$\underline{\mathbb{E}}h = \int_0^{\underline{F}(a)} h(a_{*\gamma}) d\gamma + \int_{F(a)}^{\overline{F}(a)} \min(h(a_{*\gamma}), h(a_{\gamma}^*)) d\gamma + \int_{\overline{F}(a)}^1 h(a_{\gamma}^*) d\gamma$$

shows that approximation (either outer or inner) by discretization requires at most 2 computations per discretized levels, if α is unknown.

Conditional expectations

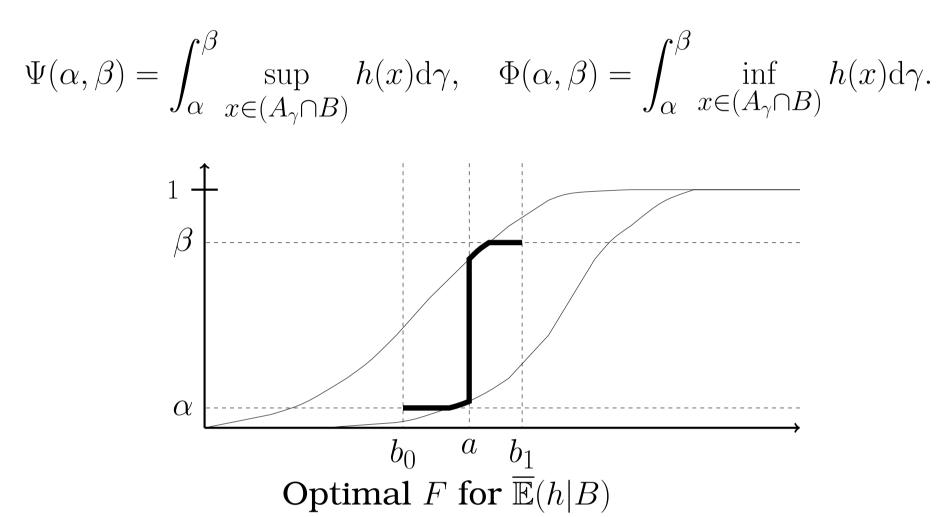
We suppose that the event $B = [b_0, b_1]$ has been observed. Lower and upper conditional expectations under B are computed as follows:

$$\underline{\mathbb{E}}(h|B) = \inf_{\underline{F} \leq F \leq \overline{F}} \frac{\int_{\mathbb{R}} h(x) I_B(x) dF}{\int_{\mathbb{R}} I_B(x) dF}, \quad \overline{\mathbb{E}}(h|B) = \sup_{\underline{F} \leq F \leq \overline{F}} \frac{\int_{\mathbb{R}} h(x) I_B(x) dF}{\int_{\mathbb{R}} I_B(x) dF}.$$

In the case of h having one maximum, these formulas become

$$\overline{\mathbb{E}}(h|B) = \sup_{\substack{\underline{F}(b_0) \leq \alpha \leq \overline{F}(b_0) \\ \underline{F}(b_1) \leq \beta \leq \overline{F}(b_1)}} \frac{1}{\beta - \alpha} \Psi(\alpha, \beta), \quad \underline{\mathbb{E}}(h|B) = \inf_{\substack{\underline{F}(b_0) \leq \alpha \leq \overline{F}(b_0) \\ \underline{F}(b_1) \leq \beta \leq \overline{F}(b_1)}} \frac{1}{\beta - \alpha} \Phi(\alpha, \beta),$$

with



- Numerator and denominator play opposite role in the evolution of expectations (e.g. for the upper one, both increase with the value of $\beta \alpha$).
- The main problem is to find the couple (α, β) for which extremal expectations are reached. One possibility is to start with $(\alpha, \beta) = (\underline{F}(b_0), \overline{F}(b_1))$ and then to shrink this interval.

Many dimensions, One global maximum

We assume h(X,Y) is a function from $\mathbb{R}^2 \to \mathbb{R}$. Our uncertainty model about X,Y becomes a bivariate p-box

$$\underline{F}(x,y) \le F(x,y) \le \overline{F}(x,y), \ \forall (x,y) \in \mathbb{R}^2.$$

- \blacktriangleright h has one global maximum at point (x_0,y_0) and is non-increasing in every direction from (x_0,y_0) .
- ➤ We study how upper/lower expectations can be computed under various assumptions of independence.

Random set corresponding to the marginal p-box of Y given by sets $B_{\kappa} = [b_{*\kappa}, b_{\kappa}^*]$ s.t.

$$b_{*\kappa} := \sup\{y \in [b_{inf}, b_{sup}] : \overline{F}(y) < \kappa\} = \overline{F}^{-1}(\kappa),$$

$$b_{\kappa}^* := \inf\{y \in [b_{inf}, b_{sup}] : \underline{F}(y) > \kappa\} = \underline{F}^{-1}(\kappa).$$

Case of strong independence (LP)

- If h separable (i.e. $h(X,Y)=h_1(X)h_2(Y)$), then under strong independence, $\underline{\mathbb{E}}(h)=\underline{\mathbb{E}}(h_1)\cdot\underline{\mathbb{E}}(h_2)$ and $\overline{\mathbb{E}}(h)=\overline{\mathbb{E}}(h_1)\cdot\overline{\mathbb{E}}(h_2)$.
- ightharpoonup If h not separable, then, under our assumptions and by LP approach, we get the formula

$$\overline{\mathbb{E}}(h(X,Y)) = \sup_{\underline{F}_2 \le F_2 \le \overline{F}_2} \int_{\mathbb{R}} \overline{\mathbb{E}}(h(X,z)) dF_2(z) = \sup_{\xi(y_0)} \xi(y_0) \left[\overline{F}_2(y_0) - \underline{F}_2(y_0) \right] + \int_{-\infty}^{y_0} \sup_{\xi(z)} \xi(z) d\underline{F}_2(z) + \int_{y_0}^{\infty} \sup_{\xi(z)} \xi(z) d\overline{F}_2(z)$$

where

$$\sup_{\underline{F}_1 \leq F_1 \leq \overline{F}_1} \xi(z) = h(x_0, z) \left[\overline{F}_1(x_0) - \underline{F}_1(x_0) \right] + \int_{-\infty}^{x_0} h(x, z) d\underline{F}_1 + \int_{x_0}^{\infty} h(x, z) d\overline{F}_1.$$

- This explicit formula comes down to concentrate probability mass on (x_0, y_0) and is similar to the one obtained for the univariate case.
- Formula obtained for lower expectation is

$$\underline{\mathbb{E}}(h(X,Y)) = \inf_{\underline{F}_2 \leq F_2 \leq \overline{F}_2} \int_{\mathbb{R}} \underline{\mathbb{E}}(h(X,z)) dF_2(z)$$

$$= \int_{-\infty}^{\overline{F}_2^{-1}(\beta)} \int_{-\infty}^{\overline{F}_1^{-1}(\alpha_z)} h(x,z) d\overline{F}_1 d\overline{F}_2 + \int_{-\infty}^{\overline{F}_2^{-1}(\beta)} \int_{\underline{F}_1^{-1}(\alpha_z)}^{\infty} h(x,z) d\underline{F}_1 d\overline{F}_2$$

$$+ \int_{F_2^{-1}(\beta)}^{\infty} \int_{-\infty}^{\overline{F}_1^{-1}(\alpha_z)} h(x,z) d\overline{F}_1 d\underline{F}_2 + \int_{F_2^{-1}(\beta)}^{\infty} \int_{F_1^{-1}(\alpha_z)}^{\infty} h(x,z) d\underline{F}_1 d\underline{F}_2.$$

where α_z is a solution of equation $h(\overline{F}_1^{-1}(\alpha), z) = h(\underline{F}_1^{-1}(\alpha), z)$ and β solution of $\underline{\mathbb{E}}(h(X, \underline{F}_2^{-1}(\beta))) = \underline{\mathbb{E}}(h(X, \overline{F}_2^{-1}(\beta)))$.

Again, "transitions" levels α_z, β have to be found, most of the time by numerical approximations.

For a n dimensional function with one global maximum, n such levels must be found to compute lower expectation.

Case of random set independence (RS)

Given marginal random sets, we have

$$\underline{\mathbb{E}}(h) = \int_0^1 \int_0^1 \inf_{(x,y) \in [B_{\kappa} \times A_{\gamma}]} h(x,y) d\kappa d\gamma, \quad \overline{\mathbb{E}}(h) = \int_0^1 \int_0^1 \sup_{(x,y) \in [B_{\kappa} \times A_{\gamma}]} h(x,y) d\kappa d\gamma,$$

- Again, solution can be (outer or inner) approximated by discretized levels, the main difficulty being to find the inf, sup (here, at most 4 computations are needed per descretized levels)
- From a numerical standpoint, RS ind. equivalent to 1^{st} order Monte-carlo sim. where A_{γ}, B_{κ} are randomly sampled.
- Interest: random set independence computationally attractive, while result is an outer approximation of results in case of strong and epistemic independence.

Case of unknown interaction

- ► Given random set marginals, unknown interaction is quivalent to consider every possible joint random sets having those for marginals.
- Method: approximate $[\underline{F}, \overline{F}]_X$, $[\underline{F}, \overline{F}]_Y$ with sets A_{γ_i} , B_{κ_j} ($i, j = 1, \ldots, n$) and where all sets have equal weights. Then compute (for an approximation of lower expectation)

$$\underline{\mathbb{E}}^*(h) = \inf_{\Gamma_{\gamma,\kappa} \in \Gamma_{\gamma,\kappa}^*} \sum_{\substack{x \in A_{\gamma_i} \\ y \in B_{\kappa,i}}} h(x,y) m_{\Gamma_{\gamma,\kappa}} (A_{\gamma_i} \times B_{\kappa_j})$$

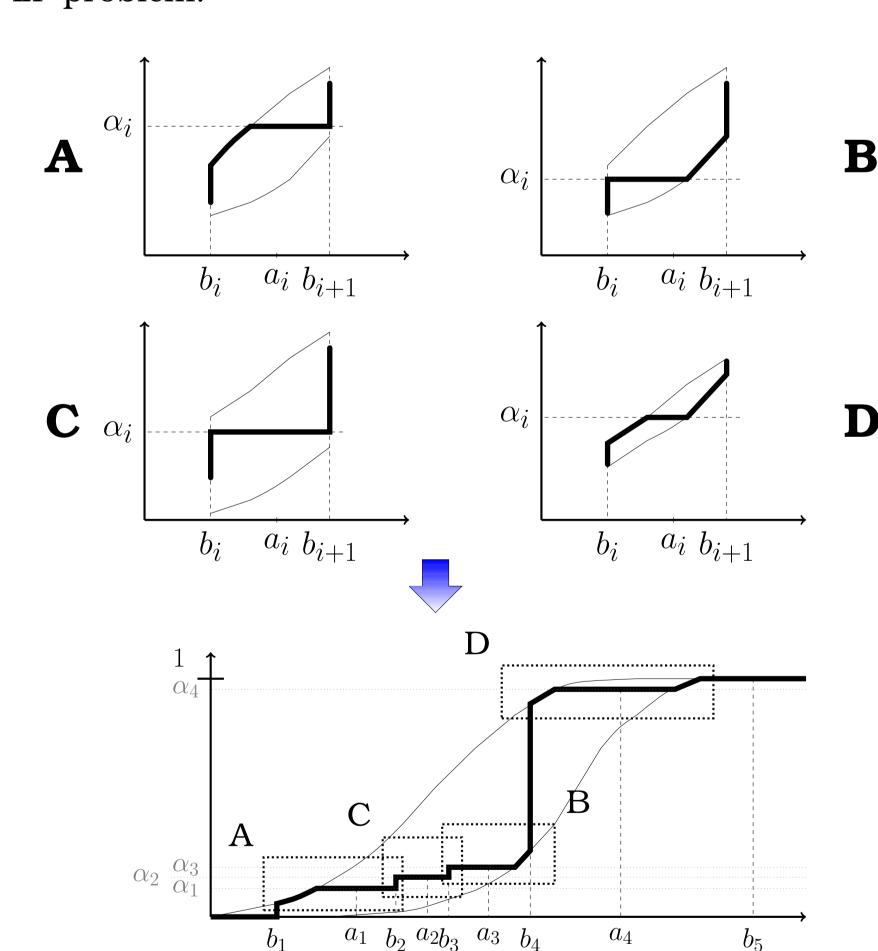
subject to

$$\sum_{j=1}^{n} m_{\Gamma_{\gamma,\kappa}}(A_{\gamma_i} \times B_{\kappa_j}) = m_{\Gamma_{\gamma}}(A_{\gamma_i}), \quad \sum_{i=1}^{n} m_{\Gamma_{\gamma,\kappa}}(A_{\gamma_i} \times B_{\kappa_j}) = m_{\Gamma_{\gamma}}(B_{\kappa_j}),$$

where $\Gamma_{\gamma,\kappa}^*$ is the set of joint random sets. $\overline{\mathbb{E}}^*(h)$ is computed by replacing inf with \sup .

One dimension, many extrema

- \blacktriangleright h has alternate local maxima (a_i) and local minima (b_i).
- LP approach shows that optimal F reaching $\underline{\mathbb{E}}(h)$ is a combination of four different local subcases that are part of a large LP problem.



Example of Optimal F with general \boldsymbol{h} which extrema are known

These four subcases can be found back in the following formula using random sets

$$\underline{\mathbb{E}}(h) = \int_0^{\underline{F}(b_n)} \min_{b_i \in A_{\gamma}} (h(a_{*\gamma}), h(b_i), h(a_{\gamma}^*)) d\gamma + \int_{F(b_n)}^1 h(a_{*\gamma}) d\gamma,$$

Optimal distribution F is a succession of vertical jumps (prob. mass concentrated on b_i) and of horizontal jumps (to avoid highest values of h)

perspectives

- Pursue investigations on multivariate case, by generalizing existing results to more general functions and to n dimensional case and by exploring the case of epistemic independence
- Design efficient algorithms to make good approximations (i.e. how to find good values for levels α_i with functions having many extrema?)
- Study various ways to integrate information about dependencies, e.g. by using copulas or adding constraints to LP problems.